

# *The behavior of the general stable distributions and their Fisher information matrix near the normal distribution*

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# Abstract

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## Abstract

- Introduction and applications.
- Expressions of the characteristic function and the density of a general stable distribution.
- The density of a general stable distribution close to the normal distribution.
- Fisher information close to the normal distribution.
- Numerical study (symmetric case).

# Heavy tailed data

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## Example

- Financial data.
- Network models.
- Telecommunication

## How to deal with ?

- Extreme value theory
- Heavy tailed distribution (e.x. pareto distribution)
- Heavy tailed time series model (e.x. GARCH, SV)
- Copulas

Note that stable distributions can model very heavy tailed data only. (no second moments)

## Definition and Properties

### Definition

- The distribution  $R$  is stable (in the broad sense) if for each  $n$  there exist constants  $c_n > 0$ ,  $\gamma_n$  such that

$$S_n \stackrel{d}{=} c_n X + \gamma_n$$

and  $R$  is not concentrated at one point.  $R$  is stable in the strict sense with  $\gamma_n = 0$ .

Widely known distributions;

Normal distribution, Cauchy distribution and Lévy distribution.

Except these 3 distributions, Probability densities have no analytic expression.

Tail is very tick;

Except Normal distribution, no second moment exists.

## Definition and Properties

### Domain of Attraction

- The distribution  $F$  of the independent random variables  $X_k$  belongs to the domain of attraction of a distribution  $R$  if there exist norming constants  $a_n > 0$ ,  $b_n$  such that the distribution of  $a_n^{-1}(S_n - b_n)$  tends to  $R$ .

A distribution  $R$  possesses a domain of attraction iff it is stable.

Many applications

(see recently published book, Uchaikin and Zolotarev (1999)):

Chaos, fractal, physics, astrophysics, cosmology.

In economics financial applications are expected.

## Characteristic function: $\Phi(t; \alpha, \beta)$

Zolotarev's (M) parameterization (see p.11 of Zolotarev (1986)).

$$\Phi(t; \mu, \sigma, \alpha, \beta) = \exp\left(-|\sigma t|^\alpha \left[1 + i\beta \left(\tan \frac{2\alpha}{\pi}\right) (\operatorname{sgn} t) (|\sigma t|^{1-\alpha} - 1)\right] + i\mu t\right).$$

$$0 < \alpha \leq 2, \quad -1 \leq \beta \leq 1, \quad -\infty < \mu < \infty, \quad \sigma > 0.$$

location:  $\mu$ , scale:  $\sigma$ , kurtosis:  $\alpha$ , skewness:  $\beta$ .

We consider the standard case  $(\mu, \sigma) = (0, 1)$ .

$$\Phi(t; \alpha, \beta) = \exp\left\{-|t|^\alpha \left[1 + i\beta (\operatorname{sgn} t) \left(\tan \frac{2\alpha}{\pi}\right) (|t|^{1-\alpha} - 1)\right]\right\}.$$

Location scale family  $f(x; \alpha, \beta) := f(x; 0, 1, \alpha, \beta)$ :

$$f(x; \mu, \sigma, \alpha, \beta) = \frac{1}{\sigma} f\left(\frac{x - \mu}{\sigma}; \alpha, \beta\right).$$

## Density: $f(x, \alpha, \beta)$

Let

$$\zeta = -\beta \tan \frac{\pi\alpha}{2},$$

$$\varrho = \frac{2}{\pi\alpha} \arctan\left(\beta \tan \frac{\pi\alpha}{2}\right).$$

From (2.2.18) of Zolotarev (1986), for  $\alpha \neq 1$  and  $x > \zeta$ ,

$$f(x; \alpha, \beta) = \frac{\alpha(x - \zeta)^{1/(\alpha-1)}}{2|\alpha - 1|} \int_{-\varrho}^1 A(\varphi) \exp\left(- (x - \zeta)^{\alpha/(\alpha-1)} A(\varphi)\right) d\varphi.$$

$$A(\varphi; \alpha, \beta) = \left(\cos \frac{\pi}{2} \alpha \varrho\right)^{\frac{1}{\alpha-1}} \left(\frac{\cos \frac{\pi}{2} \varphi}{\sin \frac{\pi}{2} \alpha(\varphi + \varrho)}\right)^{\frac{\alpha}{\alpha-1}} \frac{\cos \frac{\pi}{2} (\alpha \varrho + (\alpha - 1)\varphi)}{\cos \frac{\pi}{2} \varphi},$$

$$f(x; \alpha, \beta) = f(-x; \alpha, -\beta) \text{ for } x < \zeta.$$

## Normal: $f(x; 2)$

As  $\alpha \uparrow 2$ ,  $\zeta$  and  $\varrho \rightarrow 0$ .

Unusual representation of the normal distribution

( $\mu = 0$ ,  $\sigma = 2$ ),

$$f(x; 2) = x \int_0^1 1 / (2 \sin \frac{\pi}{2} \varphi)^2 \exp(-x^2 / (2 \sin \frac{\pi}{2} \varphi)^2) d\varphi.$$



## Definition of problem

### MLE of stable distributions

- Brorsen and Yang (1990),
- Nolan (2001),
- Matsui and Takemura (2004).

### Behavior of the Fisher information matrix

- DuMouchel (1975, 1983) proved Fisher information  $I_{\alpha\alpha}$  (w.r.t  $\alpha$ )  $\rightarrow \infty$  as  $\alpha \rightarrow 2$ .
- Nagaev and Shkol'nik (1988) derived asymptotic behavior of  $I_{\alpha\alpha}$  as  $\alpha \rightarrow 2$  in symmetric case.

$$I_{\alpha\alpha} = \frac{1}{4\Delta \log(1/\Delta)} (1 + o(1)), \quad \Delta = 2 - \alpha.$$

## Definition of problem

Nagaev and Shkol'nik (1988) stated  
“We note the problems under study are as yet unresolved for non-symmetric stable distributions.”

We clarify the limiting values of the  $4 \times 4$  Fisher Information matrix with respect to  $\mu$ ,  $\sigma$ ,  $\alpha$  and  $\beta$ .

From here let

$$\Delta = 2 - \alpha,$$

$$\beta^* = \beta \operatorname{sgn}(x - \zeta).$$

and let  $w(t) > 0$  be a function which satisfies the property

$$\lim_{t \rightarrow 0} w(t) = 0.$$

## Theorem 2.1: Density near the normal distribution

Let  $|\beta| \neq 1$  be fixed. We define

$$\begin{aligned}F_1(x; \alpha, \beta) &= f(|x - \zeta|; 2), \\F_2(x; \alpha, \beta) &= \Delta(1 + \beta^*)|x - \zeta|^{\Delta-3}, \\g(x; \alpha, \beta) &= F_1(x; \alpha, \beta) + F_2(x; \alpha, \beta).\end{aligned}$$

Then for an arbitrarily small  $\epsilon > 0$  there exist  $\Delta_0$  and  $x_0$  such that for all  $\Delta < \Delta_0$  and  $|x| > x_0$ ,

$$|f(x; \alpha, \beta)/g(x; \alpha, \beta) - 1| < \epsilon.$$

Furthermore, for an arbitrarily small constant  $\delta > 0$ ,

$$g(x; \alpha, \beta) = \begin{cases} F_1(x; \alpha, \beta)(1 + w(\Delta)) & \text{if } |x - \zeta| \leq (2 - \delta)(\log 1/\Delta)^{1/2} \\ F_2(x; \alpha, \beta)(1 + w(\Delta)) & \text{if } |x - \zeta| \geq (2 + \delta)(\log 1/\Delta)^{1/2}. \end{cases}$$

# Visualization: Integrand near Normal

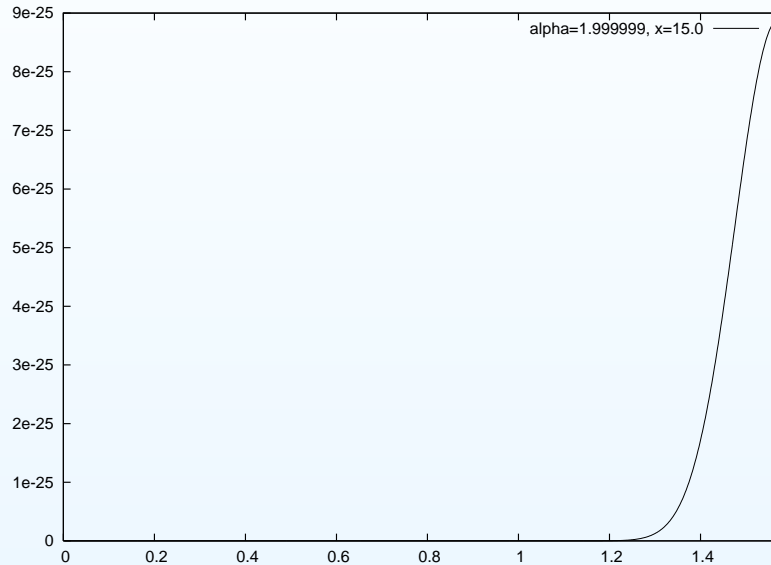


Figure 1:  $\alpha = 1.999999$ ,  $x = 15.0$

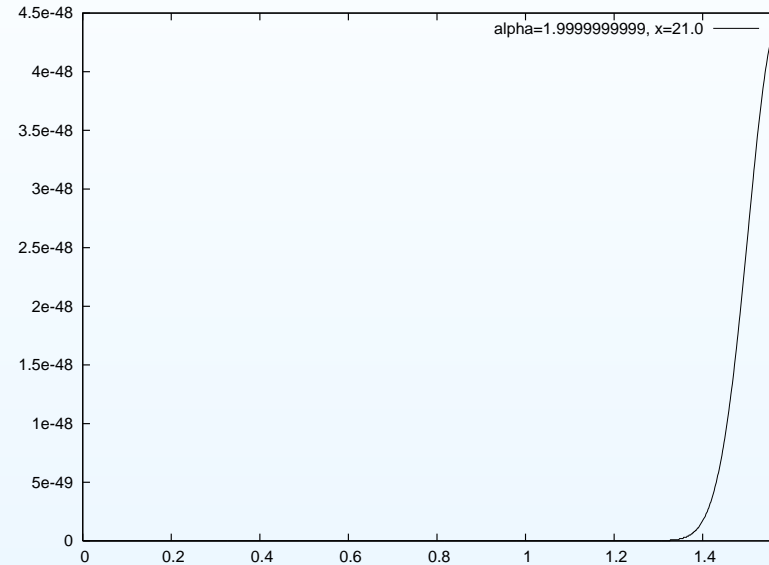


Figure 2:  $\alpha = 1.9999999999$ ,  $x = 21.0$

The integrand of density near Normal and  $x \rightarrow \infty$ .

$$f(x; \alpha, \beta) = \frac{\alpha(x - \zeta)^{1/(\alpha-1)}}{2|\alpha - 1|} \int_{-e}^1 A(\varphi) \exp\left(- (x - \zeta)^{\alpha/(\alpha-1)} A(\varphi)\right) d\varphi.$$

## Proof

Assume  $x - \zeta > 0$ .

Notations:

$\epsilon, \epsilon', \tau, \eta$  := arbitrarily small positive numbers,

$c$  := any positive constant,

$\lambda$  :=  $1 - \varphi$ ,

$\varphi_{\Delta}$  :=  $1 - \Delta^{1/2-\epsilon}$ ,

$z$  :=  $(x - \zeta)^{\alpha/(\alpha-1)} := x^2$ ,

$\varphi_0$  :=  $\varphi_{\Delta} - z^{-1/2+\tau}$ .

$$f(x; \alpha, \beta) = \frac{\alpha(x - \zeta)^{1/(\alpha-1)}}{2|\alpha - 1|} \times H.$$

## Proof

We divide integral

$$H = \int_{-\varrho}^1 A(\varphi; \alpha, \beta) \exp(-z A(\varphi; \alpha, \beta)) d\varphi,$$

into

$$H = \sum_{k=1}^6 H_k$$

where each  $H_k$  corresponds to the integration of  $H$  for  $k$ -th intervals of

$$\begin{aligned} &[-\varrho, 1 - \mu), \quad [1 - \mu, \varphi_0), \quad [\varphi_0, \varphi_\Delta) \\ &[\varphi_\Delta, 1 - \Delta/\epsilon'), \quad [1 - \Delta/\epsilon', 1 - \Delta\epsilon'), \quad [1 - \Delta\epsilon', 1]. \end{aligned}$$

## Lemmas

**Lemma 2.1 :**  $A(\varphi)$  on  $[1 - \Delta/\epsilon', 1 - \Delta\epsilon')$

As  $\Delta \rightarrow 0$  and for  $0 \leq \lambda \leq \Delta/\epsilon'$ ,

$$A(1 - \lambda) = \frac{(\lambda/\Delta)^{\frac{1}{1-\Delta}} (1 + \beta + \lambda/\Delta)}{(1 + \beta + 2\lambda/\Delta)^2} (1 + o(\Delta \log(1/\Delta))).$$

**Lemma 2.2 :**  $A(\varphi)$  on  $[\varphi_0, \varphi_\Delta)$  and  $[\varphi_\Delta, 1 - \Delta/\epsilon')$

As  $\Delta \rightarrow 0$

$$A(\varphi_\Delta) = \frac{1}{4} + \frac{\pi^2}{16} \Delta^{1-2\epsilon} + o(\Delta),$$

$$A'(\varphi_\Delta) = -\frac{\pi^2}{8} \Delta^{1/2-\epsilon} - \frac{(1 + \beta)^2}{8} \Delta^{1/2+3\epsilon} + o(\Delta).$$

## Lemmas

**Lemma 2.3:**  $A(\varphi)$  on  $[\varphi_0, \varphi_\Delta)$

As  $\Delta \rightarrow 0$  and  $\lambda \rightarrow 0$

for  $\varphi = 1 - \lambda < \varphi_\Delta = 1 - \Delta^{1/2-\epsilon}$ ,

$$A(1 - \lambda) = \frac{1}{4} + \frac{\pi^2}{16}\lambda^2 + o(\lambda^2),$$

$$A'(1 - \lambda) = -\frac{\pi^2}{8}\lambda - \frac{\Delta^2(1 - \lambda + \beta)^2}{8\lambda^3} + o(\lambda^2),$$

$$\begin{aligned} A''(1 - \lambda) &= \frac{\pi^2}{8} + \frac{\pi^2}{8}\lambda^2 - \frac{3\Delta^2(1 - \lambda + \beta)^2}{8\lambda^4} \\ &+ \frac{3\Delta^2(1 - \lambda + \beta)}{4\lambda^3} \left( 1 - \frac{\Delta(1 - \lambda + \beta)^2}{\lambda^2} \right) \\ &+ o(\lambda^2). \end{aligned}$$



## Lemmas

Note since

$0 \leq \varphi \leq \varphi_{\Delta} \Leftrightarrow \Delta/\epsilon \leq \Delta^{1/2-\epsilon} \leq \lambda \leq 1 \Rightarrow \Delta = o(\lambda^2)$ ,  
we have to consider terms like  $\Delta^2/\lambda^4$  or  $\Delta^2/\lambda^3$  in Lemma 2.3.

### **Lemma 2.4 :**

If  $\Delta$  is sufficiently small,  $A(\varphi)$  is a monotonically decreasing function on  $(-\varrho, 1)$  (integral range).

## Calculation of $H_1$

Lemma 2.1  $\Rightarrow A(\varphi_\Delta) > \frac{1}{4}$ .

Lemma 2.4 and  $1 - \mu < \varphi_\Delta$

$\Rightarrow \exists \rho \in (0, 1)$  such that  $\rho A(1 - \mu) > \frac{1}{4}$ .

$$\begin{aligned} H_1 &= \frac{1}{z(1 - \rho)} \int_{-\varrho^*}^{1 - \mu} z(1 - \rho) A(\varphi) \exp\{-z\rho A(\varphi) - z(1 - \rho) A(\varphi)\} d\varphi \\ &\leq \frac{1}{z(1 - \rho)} \exp(-z\rho A(1 - \mu)) \\ &= O(\exp(-\gamma z)/z), \quad \text{for } \gamma > \frac{1}{4}. \end{aligned}$$

## Calculation of $H_2$

Lemma 4  $\Rightarrow$

$$\begin{aligned} H_2 &= \int_{1-\mu}^{\varphi_0} A(\varphi) \exp(-zA(\varphi)) d\varphi \\ &= \mu A(1-\mu) \exp(-zA(\varphi_0)) \\ &= O(\exp(-zA(\varphi_0))). \end{aligned}$$

$$A(\varphi_0) = A(\varphi_\Delta) + A'(\varphi_\Delta)(\varphi_0 - \varphi_\Delta) + \frac{1}{2}A''(\xi)(\varphi_0 - \varphi_\Delta)^2,$$

Lemma 2.2, Lemma 2.3 and  $\varphi_0 - \varphi_\Delta = -z^{-1/2+\tau} \Rightarrow$

$$H_2 = O\left(\exp\left(-\frac{z}{4} - R_1(\Delta, z, \epsilon) - R_2(\Delta, z, \epsilon) - cz^{2\tau}\right)\right),$$

$$R_1(\Delta, z, \epsilon) = O(z\Delta^{1-2\epsilon}) > 0, \quad R_2(\Delta, z, \epsilon) = O(\Delta^{1/2-\epsilon}z^{1/2+\tau}) > 0.$$

## Calculation of $H_3$

$$H_3 = \exp(-zA(\varphi_\Delta)) \times \int_{\varphi_0}^{\varphi_\Delta} A(\varphi) \exp\left(-zA'(\varphi_\Delta)(\varphi - \varphi_\Delta) - \frac{z}{2}A''(\xi)(\varphi - \varphi_\Delta)^2\right) d\varphi.$$

Lemma 2.2 and Lemma 2.3  $\Rightarrow$

$$A(\varphi) = \frac{1}{4} + O(\lambda^2),$$

$$zA(\varphi_\Delta) = \frac{z}{4} + R_1(\Delta, z, \epsilon),$$

$$zA'(\varphi_\Delta)(\varphi - \varphi_\Delta) = R_3(\Delta, z, \epsilon), \quad 0 < R_3(\Delta, z, \epsilon) \leq R_2(\Delta, z, \epsilon),$$

$$\frac{z}{2}A''(\xi)(\varphi - \varphi_\Delta)^2 = \frac{z}{2} \frac{\pi^2}{8} (\varphi - \varphi_0)^2 + O(z^{-1+4\tau}) + O(\Delta^{4\epsilon} z^{2\tau}).$$

## Calculation of $H_3$

$$H_3 = R(\lambda, \Delta, z, \epsilon) \frac{1}{4} \exp\left(-\frac{z}{4}\right) \int_{\varphi_0}^{\varphi_\Delta} \exp\left(-\frac{z}{2} \frac{\pi^2}{8} (\varphi - \varphi_\Delta)^2\right) d\varphi,$$

$$R(\lambda, \Delta, z, \epsilon) = (1 + o(1)) \exp(-R_1(\Delta, z, \epsilon) - R_3(\Delta, z, \epsilon)).$$

$$\begin{aligned} & \int_{\varphi_0}^{\varphi_\Delta} \exp\left(-\frac{z}{2} \frac{\pi^2}{8} (\varphi - \varphi_0)^2\right) d\varphi \\ &= \frac{1}{\sqrt{z}} \int_0^\infty \exp\left(-\frac{z}{2} \frac{\pi^2}{8} \varphi^2\right) d\varphi - \frac{1}{\sqrt{z}} \int_{z^\tau}^\infty \exp\left(-\frac{z}{2} \frac{\pi^2}{8} \varphi^2\right) d\varphi \\ &= \frac{2}{\sqrt{\pi z}} - O\left(\frac{1}{z^{1/2+\tau}} \exp\left(\frac{-z^{2\tau}}{2}\right)\right). \end{aligned}$$

$$H_3 = \frac{1}{2\sqrt{\pi z}} \exp\left(-\frac{z}{4} - R_1(\Delta, z, \epsilon) - R_3(\Delta, z, \epsilon)\right) (1 + o(1)).$$

## Calculation of $H_4$

Lemma 2.4  $\Rightarrow$

$$\begin{aligned} H_4 &\leq \int_{\varphi_\Delta}^{1-\Delta/\epsilon'} A(\varphi_\Delta) \exp(-zA(1 - \Delta/\epsilon')) d\varphi \\ &\leq \Delta^{1/2-\epsilon} A(\varphi_\Delta) \exp(-zA(1 - \Delta/\epsilon')). \end{aligned}$$

$\lambda/\Delta \leq 1/\epsilon'$ , Lemma 2.1 and 2.4  $\Rightarrow$

$$A(1 - \Delta/\epsilon') = \frac{1}{4}(1 - w(\epsilon')), \quad \text{as } \Delta \rightarrow 0.$$

$\Rightarrow$

$$H_4 \leq O \left( \Delta^{1/2-\epsilon} \exp \left( -\frac{z}{4}(1 - w(\epsilon')) \right) \right).$$

## Calculation of $H_5$

From Lemma 1 for  $\epsilon' \leq \lambda/\Delta \leq 1/\epsilon'$ ,  $A(\varphi)$  is bounded. Then

$$\begin{aligned} H_5 &= \int_{1-\Delta/\epsilon'}^{1-\Delta\epsilon'} A(\varphi) \exp(-zA(\varphi)) d\varphi \\ &\leq (\Delta/\epsilon' - \Delta\epsilon') \sup_{\varphi \in (1-\Delta/\epsilon', 1-\Delta\epsilon')} A(\varphi) \exp(-zA(\varphi)) \\ &\leq O(\Delta \exp(-cz)). \end{aligned}$$

## Calculation of $H_6$

Lemma 2.1  $\Rightarrow$  for  $\lambda/\Delta \leq \epsilon'$

$$A(\varphi; \alpha, \beta) = \frac{1 + R_4(\Delta, \epsilon')}{1 + \beta} (\lambda/\Delta)^{1/(1-\Delta)},$$

$$R_4(\Delta, \epsilon') = O(\epsilon') + O(\Delta \log(1/\Delta)).$$

$\varphi = 1 - \lambda \rightarrow \lambda$  in  $H_6$

$$H_6 = \left( \frac{1 + R_4}{1 + \beta} \right) \int_0^{\Delta \epsilon'} (\lambda/\Delta)^{1/(1-\Delta)} \exp \left( -z \left( \frac{1 + R_4}{1 + \beta} \right) (\lambda/\Delta)^{1/(1-\Delta)} \right) d\lambda.$$

$$\lambda = \Delta \left( \frac{1 + \beta}{1 + R_4} \frac{x}{z} \right)^{1-\Delta} \rightarrow x, \quad g(z) = \epsilon'^{1/(1-\Delta)} \left( \frac{1 + R_4}{1 + \beta} \right) z,$$

$$H_6 = \frac{\Delta(1 + \beta)^{1-\Delta}}{z^{2-\Delta}} \int_0^{g(z)} x \exp(-x) dx \times (1 + o(1)).$$



## Calculation of $H_6$

$$\begin{aligned}\int_0^{g(z)} x \exp(-x) dx &= \left[ -xe^{-x} - e^{-x} \right]_0^{g(z)} \\ &= 1 + O(ge^{-g}),\end{aligned}$$

$g(z) \rightarrow \infty$  as  $z \rightarrow \infty$ . Therefore

$$H_6 = \Delta(1 + \beta)z^{\Delta-2}(1 + o(1)).$$

## Review

$$f(x; \alpha, \beta) = \frac{\alpha(x - \zeta)^{1/(\alpha-1)}}{2|\alpha - 1|} \times H.$$

$$H = \int_{-\varrho}^1 A(\varphi; \alpha, \beta) \exp(-z A(\varphi; \alpha, \beta)) d\varphi,$$

$$H = \sum_{k=1}^6 H_k$$

where each  $H_k$  corresponds to the integration of  $H$  for  $k$ -th intervals of

$$\begin{aligned} &[-\varrho, 1 - \mu), \quad [1 - \mu, \varphi_0), \quad [\varphi_0, \varphi_\Delta) \\ &[\varphi_\Delta, 1 - \Delta/\epsilon'), \quad [1 - \Delta/\epsilon', 1 - \Delta\epsilon'), \quad [1 - \Delta\epsilon', 1]. \end{aligned}$$

## Calculation of $H$

$$H_1 = O(\exp(-\gamma z)/z), \quad \text{for } \gamma > \frac{1}{4}$$

$$H_2 = O\left(\exp\left(-\frac{z}{4} - R_1(\Delta, z, \epsilon) - R_2(\Delta, z, \epsilon) - cz^{2\tau}\right)\right)$$

$$H_3 = \frac{1}{2\sqrt{\pi z}} \exp\left(-\frac{z}{4} - R_1(\Delta, z, \epsilon) - R_3(\Delta, z, \epsilon)\right) (1 + o(1)).$$

$$H_4 \leq O\left(\Delta^{1/2-\epsilon} \exp\left(-\frac{z}{4}(1 - w(\epsilon'))\right)\right)$$

$$H_5 \leq O(\Delta \exp(-cz))$$

$$H_6 = \Delta(1 + \beta)z^{\Delta-2}(1 + o(1)).$$

$\Rightarrow$

$$\max(H_1, H_2, H_4, H_5) = o(\max(H_3, H_6)).$$

## Calculation of $H$

For an arbitrarily small  $\delta > 0$ ,

$$H_6 = o(H_3) \quad \text{if } z \leq (4 - \delta) \log 1/\Delta,$$

$$H_3 = o(H_6) \quad \text{if } z \geq (4 + \delta) \log 1/\Delta.$$

If  $z \leq (4 - \delta) \log 1/\Delta$ ,  $R_1(\Delta, z, \epsilon) \downarrow 0$ ,  $R_3(\Delta, z, \epsilon) \downarrow 0$ ,

$$H_3 = \frac{1}{2\sqrt{\pi z}} \exp\left(-\frac{z}{4}\right) (1 + o(1)).$$

□

## $f_\alpha(x; \alpha, \beta)$

### Lemma 3.1:

As  $\Delta = 2 - \alpha \rightarrow 0$ , there exists  $x_0$  and for all  $|x| \geq x_0$ ,

$$f_\alpha(x; \alpha, \beta) = -\frac{1}{|y|^{1+\alpha}} \left( 1 + \beta^* + \Delta(M_1 + M_2 \log |y|) + \frac{M_3}{|y|} + \frac{M_4 + M_5 \log |y|}{|y|^\alpha} \right) + o\left(\frac{1}{|y|^{1+2\alpha}}\right),$$

where  $y = x - \zeta$  and  $M_i$ 's are some constants.

□

Lemma 3.1 covers the case of  $\alpha = 2$ .

# Information matrix

## Theorem 4.1:

As  $\Delta = 2 - \alpha \rightarrow 0$ , Behavior of Fisher information matrix of general stable distributions when  $\mu = 0, \sigma = 1, \beta \neq \pm 1$  is given as follows.

$$\begin{bmatrix} I_{\mu\mu} & I_{\mu\sigma} & I_{\mu\alpha} & I_{\mu\beta} \\ * & I_{\sigma\sigma} & I_{\sigma\alpha} & I_{\sigma\beta} \\ * & * & I_{\alpha\alpha} & I_{\alpha\beta} \\ * & * & * & I_{\beta\beta} \end{bmatrix} = \begin{bmatrix} 0.5 + o(1) & o(1) & o(1) & O(\Delta) \\ * & 2.0 + o(1) & -\frac{1}{2} \log \log 1/\Delta & o(\Delta \log \log 1/\Delta) \\ * & * & \frac{1}{4\Delta \log 1/\Delta} & o\left(\frac{1}{\log 1/\Delta}\right) \\ * & * & * & \frac{\Delta}{4(1-\beta^2) \log 1/\Delta} \end{bmatrix}.$$

## Information matrix

Table 1: Limit of information matrix at  $\alpha = 2$

$I_{\theta\theta}$	$\mu$	$\sigma$	$\alpha$	$\beta$
$\mu$	0.5	0	0	0
$\sigma$	0	2.0	$-\infty$	0
$\alpha$	0	$-\infty$	$\infty$	0
$\beta$	0	0	0	0

# Visualization: Density and Score in Symmetric Case

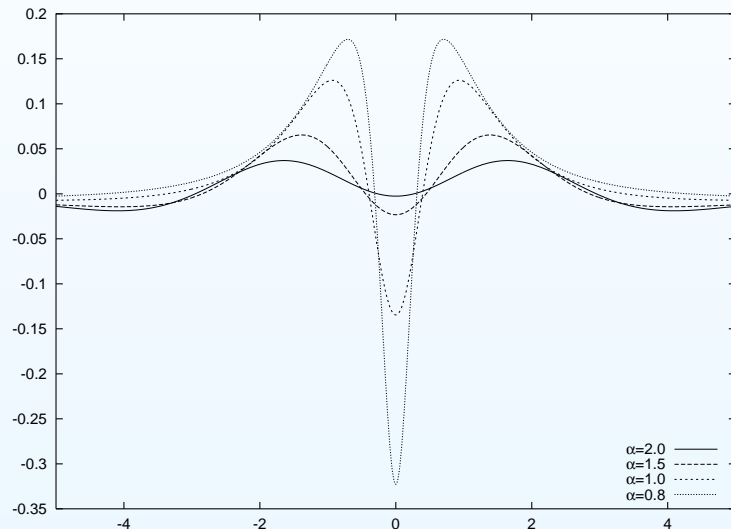


Figure 3:  $f_\alpha(x; \alpha, 0)$

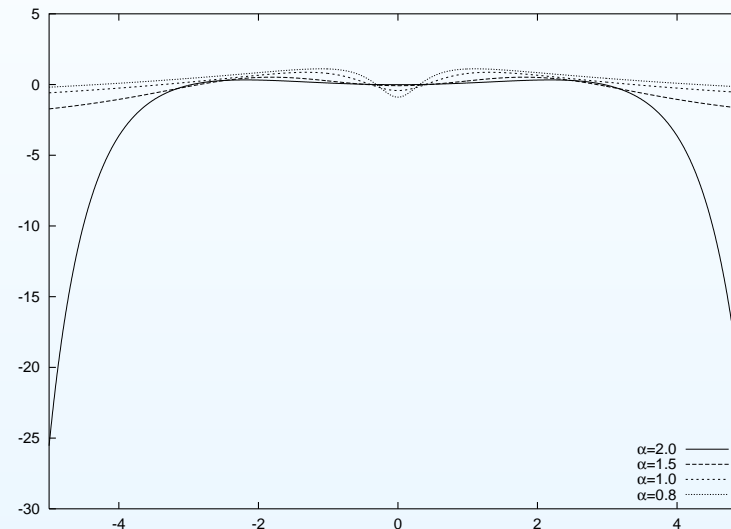


Figure 4:  $f_\alpha(x; \alpha, 0) / f(x; \alpha, 0)$

While  $f(x; \alpha, 0)$  is exponential order as  $x \rightarrow \infty$  at  $\alpha = 2$ ,  
 $f_\alpha(x; \alpha, 0)$  is polynomial order as  $x \rightarrow \infty$  at  $\alpha = 2$ .  
Therefore Fisher information  $I_{\alpha\alpha} \uparrow \infty$ .



# Numerical work

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## Numerical work

- Maximum likelihood estimation.
- Observed Fisher information matrix.
- Numerical confirmation of  $I_{\alpha\alpha}$  as  $\alpha \rightarrow 2$ .

All numerical works are done in symmetric case.

## Maximum Likelihood Estimation (Symmetric Case)

n	$\hat{\mu}$	$\hat{\sigma}$	$\hat{\alpha}$	$\hat{I}_{\mu\mu}$	$\hat{I}_{\sigma\sigma}$	$\hat{I}_{\alpha\alpha}$
	0	1.0	2.0	0.5	2.0	$\infty$
50	0.00014	0.977	1.976	0.607	1.875	5.647
100	0.00017	0.975	1.990	0.868	1.402	9.445
200	0.00094	0.977	1.994	0.784	1.239	12.77
	0	1.0	1.8	0.4552	1.3898	0.5937
50	0.0012	0.991	1.818	0.487	1.231	0.676
100	0.0033	1.002	1.822	0.482	1.356	0.584
200	-0.0000	1.000	1.810	0.450	1.399	0.603
	0	1.0	1.5	0.4281	0.9556	0.4737
50	-0.0022	1.012	1.548	0.3161	0.5796	0.4252
100	-0.0000	1.000	1.524	0.3914	0.9138	0.4278
200	-0.0003	1.000	1.510	0.4028	0.9474	0.4683
	0	1.0	1.0	0.5	0.5	0.8590
50	-0.0029	0.996	1.026	0.4243	0.4877	0.670
100	-0.0041	0.988	1.001	0.4438	0.4205	0.746
200	-0.0039	1.001	1.006	0.4929	0.5013	0.845
	0	1.0	0.8	0.6800	0.3586	1.3928
50	-0.0016	1.005	0.815	0.5434	0.3243	1.111
100	-0.0001	1.003	0.811	0.6015	0.3459	1.171
200	-0.0009	1.000	0.805	0.6232	0.3708	1.303

## Observed Fisher information

Observed Fisher information w.r.t.  $\alpha$

$$\begin{aligned}\hat{I}_{\alpha\alpha}(x_1, \dots, x_n) &= -\frac{1}{n} \sum_{i=1}^n \frac{\partial^2 \log f(x_i; \hat{\alpha})}{\partial \alpha^2} \\ &= \frac{1}{n} \sum_{i=1}^n \left( \frac{f_{\alpha}(x_i; \hat{\alpha})}{f(x_i; \hat{\alpha})} \right)^2 - \frac{1}{n} \sum_{i=1}^n \frac{f_{\alpha\alpha}(x_i; \hat{\alpha})}{f(x_i; \hat{\alpha})}\end{aligned}$$

We define this as  $\hat{I}_{\alpha\alpha}(2)$ . Observed Fisher information only use the first derivatives

$$\hat{I}_{\alpha\alpha}(1) = \frac{1}{n} \sum_{i=1}^n \left( \frac{f_{\alpha}(x_i; \hat{\alpha})}{f(x_i; \hat{\alpha})} \right)^2$$

## Observed Fisher information

Simulated observed Fisher information ( Sample size  $n = 50$ , 1000 iterations)

In table,  $\hat{I}_{\alpha\alpha}(i)$ ,  $i = 1, 2$  are means of 1000 iterations. and the variances of  $\hat{I}_{\alpha\alpha}(i)$ ,  $i = 1, 2$  are in ().

$\alpha$	$\bar{\alpha}$	$\hat{I}_{\alpha\alpha}(1)$	$\hat{I}_{\alpha\alpha}(2)$	$I_{\alpha\alpha}$
1.5	1.531	0.4863 (0.033)	0.5145 (0.018)	0.4737
1.0	1.025	0.8541 (0.157)	0.9001 (0.099)	0.8590
0.5	0.509	4.2819 (3.517)	4.4660 (2.755)	4.2748

$\hat{I}_{\alpha\alpha}(1)$ : small bias and large variance.

$\hat{I}_{\alpha\alpha}(2)$ : small variance and large bias.

Sometimes  $\hat{I}_{\alpha\alpha}(1)$  (only use the first derivatives) is useful.

## Numerical confirmation of $I_{\alpha\alpha}$ as $\alpha \rightarrow 2$

In symmetric case we numerically examine Nagaev and Shkol'nik (1988) results.

$$I_{\alpha\alpha} = \frac{1}{4\Delta \log(1/\Delta)} (1 + o(1)), \quad \Delta = 2 - \alpha.$$

$\alpha$	$I_{\alpha\alpha}(1)$	$I_{\alpha\alpha}(2)$	N&S	$I_{\sigma\alpha}(1)$	$I_{\sigma\alpha}(2)$
$2.0 - 10^{-10}$	106860414	92384764	108573620	-1.3482	-1.3427
$2.0 - 10^{-9}$	10810787	10167389	12063736	-1.3482	-1.3395
$2.0 - 10^{-8}$	1144778	1131645	1357170	-1.3482	-1.3123
$2.0 - 10^{-7}$	127953	127802	155105.2	-1.3478	-1.2553
$2.0 - 10^{-6}$	14724	14722	18095.60	-1.2094	-1.1923
1.99999	1750	1750	2171.472	-1.1100	-1.1100
1.9999	217	217	271.4341	-1.0069	-1.0069

$I_{\alpha\alpha}(1)$  near  $x \rightarrow \infty$  Taylor expansion at  $\alpha = 2$  w.r.t.  $\alpha$ .  
 $I_{\alpha\alpha}(2)$  near  $x \rightarrow \infty$  use asymptotic expansion.

## Future work

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### What can do ? with stable

- Very heavy tail model.
- Distributions with finite order moments.

### Future work

- Tail is not so heavy (have second moment)
- Non i.i.d. dependent data.
- Multivariate data.

Copulas ? Resampling method?

important = why, for what purpose, how you use this model.

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## Proof of $I_{\alpha\alpha}$

We divide integral of Fisher information matrix into three parts.

$$\begin{aligned} I_{\alpha\alpha} &= \int_0^{\infty} \frac{\{f_{\alpha}(x + \zeta; \alpha, \beta)\}^2}{f(x + \zeta; \alpha, \beta)} dx + \int_0^{\infty} \frac{\{f_{\alpha}(x + \zeta; \alpha, -\beta)\}^2}{f(x + \zeta; \alpha, -\beta)} dx \\ &\quad + \int_{-\zeta}^{\zeta} \frac{\{f_{\alpha}(x; \alpha, \beta)\}}{f(x; \alpha, \beta)} dx \\ &= I_{\alpha\alpha}^1 + I_{\alpha\alpha}^2 + I_{\alpha\alpha}^3. \end{aligned}$$

This is obtained by the relation  $f(x; \alpha, \beta) = f(-x; \alpha, -\beta)$  for  $x - \zeta < 0$ .

## Proof of $I_{\alpha\alpha}$

$I_{\alpha\alpha}^1$  into five subintegrals,

$$I_{\alpha\alpha}^1 = \sum_{k=1}^5 I_{\alpha\alpha}(k)$$

where each  $I_{\alpha\alpha}(k)$  corresponds to the  $k$ -th intervals of  $[0, T)$ ,  $[T, x_1(\Delta))$ ,  $[x_1(\Delta), x_2(\Delta))$ ,  $[x_2(\Delta), x_3(\Delta))$ ,  $[x_3(\Delta), \infty)$ .

$$x_1(\Delta) = (2 - \delta)(\log 1/\Delta)^{1/2}, \quad x_2(\Delta) = (2 + \delta)(\log 1/\Delta)^{1/2}, \\ x_3(\Delta) = \exp(\Delta^{-1/2}).$$

**Theorem 2.1** and **Lemma 3.1**

$$I_{\alpha\alpha}(1) < \infty.$$

## Proof of $I_{\alpha\alpha}$

For  $I_{\alpha\alpha}(2)$ ,

$$f(x+\zeta; \alpha, \beta) = f(x; 2)(1+o(1)), \quad f_{\alpha}(x+\zeta; \alpha, \beta) = \text{const} \times x^{\Delta-3}.$$

$$\begin{aligned} I_{\alpha\alpha}(2) &= \text{const} \times \int_T^{x_1(\Delta)} x^{2\Delta-6} \exp\left(\frac{x^2}{4}\right) dx \\ &= \text{const} \times \frac{1}{\Delta^{1-\delta} (\log 1/\Delta)^{5/2-\Delta}}. \end{aligned}$$

For  $I_{\alpha\alpha}(3)$ ,

$$f(x+\zeta; \beta, \alpha) = \text{const} \times \Delta x^{\Delta-3}, \quad f_{\alpha}(x+\zeta; \alpha, \beta) = \text{const} \times x^{\Delta-3}.$$

$$\begin{aligned} I_{\alpha\alpha}(3) &= \text{const} \times \frac{1}{\Delta} \int_{x_1(\Delta)}^{x_2(\Delta)} x^{\Delta-3} dx \\ &= \text{const} \times \frac{\delta}{\Delta \log 1/\Delta}. \end{aligned}$$

## Proof of $I_{\alpha\alpha}$

For  $I_{\alpha\alpha}(4)$ ,

$$f(x + \zeta; \alpha, \beta) = (1 + \beta)\Delta x^{\Delta-3}(1 + o(1)),$$

$$f(x + \zeta; \alpha, \beta) = -(1 + \beta)x^{\Delta-3}(1 + o(1)).$$

$$\begin{aligned} I_{\alpha\alpha}(4) &= \frac{1 + \beta}{\Delta}(1 + o(1)) \int_{x_2(\Delta)}^{x_3(\Delta)} x^{\Delta-3} dx \\ &= \frac{1 + \beta}{8\Delta \log 1/\Delta}(1 + o(1)). \end{aligned}$$

For  $I_{\alpha\alpha}(5)$ ,

$$f(x + \zeta; \alpha, \beta) = (1 + \beta)\Delta x^{\Delta-3}(1 + o(1)),$$

$$f_{\alpha}(x + \zeta; \alpha, \beta) = -(1 + \beta + \Delta \log x)x^{\Delta-3}(1 + o(1)).$$

## Proof of $I_{\alpha\alpha}$

$$\begin{aligned} I_{\alpha\alpha}(5) &= \text{const} \times \frac{1}{\Delta} \int_{x_3(\Delta)}^{\infty} x^{\Delta-3} \{\max(1 + \beta, \Delta \log x)\}^2 dx \\ &\leq \text{const} \times \Delta \int_{x_3(\Delta)}^{e^{(1+\beta)/\Delta}} x^{\Delta-3} (\log x)^2 dx \\ &\quad + \text{const} \times \frac{1}{\Delta} \int_{e^{(1+\beta)/\Delta}}^{\infty} x^{\Delta-3} dx \\ &\leq \text{const} \times \frac{1}{\Delta} \int_{x_3(\Delta)}^{\infty} x^{\Delta-3} dx \\ &= O(e^{-2/\Delta^{1/2}} / \Delta) \rightarrow 0, \text{ as } \Delta \rightarrow 0. \end{aligned}$$

## Proof of $I_{\alpha\alpha}$

$$I_{\alpha\alpha}^1 = \frac{1 + \beta}{8\Delta \log 1/\Delta} (1 + o(1)).$$

Setting  $\beta \rightarrow -\beta$  in  $I_{\alpha\alpha}^1$ , we obtain  $I_{\alpha\alpha}^2$ . By  $\zeta = O(\Delta)$  and finiteness of the integrand of  $I_{\alpha\alpha}^3$ ,  $I_{\alpha\alpha}^3 = O(\Delta)$ .

□